Tradeoffs Between Impulse Response Inference Methods

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January 7, 2023

This talk

• Selective summary (manifesto?) of recent research about estimation/inference for

$$\theta_h \equiv E[y_{t+h} \mid x_t = 1] - E[y_{t+h} \mid x_t = 0], \quad h = 0, 1, 2, \dots$$

- Comparison of semi-structural estimation procedures:
 - Local projections (LP), vector autoregressions (VAR), shrinkage variants (Bayesian, smoothing).
- Large-sample equivalence, but finite-sample trade-offs:
 - Point estimation: bias vs. variance.
 - Inference: different notions of confidence interval coverage.
- Point out areas for future research.

- 2 Large-sample estimand
- **3** Finite-sample estimation
- 4 Confidence intervals
- Sonlinearities
- **6** Topics for future research

Assumptions

- For simplicity, focus on simple case where x_t (shock of interest) is...
 - 1 observed.
 - 2 independent of all past data.
 - **3** mean = 0, variance = 1.
- Denote the full data vector by $Y_t = (x_t, y_t, \dots)'$. Assume stationary.
- Abstract from any deterministic terms.
- These assumptions can be relaxed.

Impulse responses estimators

• VAR estimator: extrapolates $\hat{\theta}_h$ from first p sample autocov's. Sims (1980)

1 Estimate model

$$Y_t = \sum_{\ell=1}^p \hat{A}_\ell Y_{t-\ell} + \hat{u}_t, \quad \hat{\Sigma} = rac{1}{T} \sum_{t=1}^T \hat{u}_t \hat{u}_t'.$$

2 $\hat{\theta}_h = e'_2 \hat{\Psi}_h \hat{\nu}$, where $\hat{\nu} = \text{chol}(\hat{\Sigma}) e_1$, and we iterate

$$\hat{\Psi}_{h} = \sum_{\ell=1}^{\min\{p,h\}} \hat{A}_{\ell} \hat{\Psi}_{h-\ell}, \ h = 1, 2, \dots, \quad \hat{\Psi}_{0} = I.$$

• LP estimator: direct projection, separately for each h. Jordà (2005)

$$y_{t+h} = \hat{\theta}_h x_t + \sum_{\ell=1}^p \hat{\delta}'_{\ell,h} Y_{t-\ell} + \hat{\mathbf{e}}_{t,h}.$$

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LP and VAR estimate the same impulse responses

• In population, LP and VAR estimate the same impulse responses:

$$\underset{T \to \infty}{\text{plim}} \hat{\theta}_{h}^{\text{LP}} = \underset{T \to \infty}{\text{plim}} \hat{\theta}_{h}^{\text{VAR}} = \text{Proj}[y_{t+h} \mid x_t = 1] - \text{Proj}[y_{t+h} \mid x_t = 0] \equiv \theta_{h}$$

for all horizons $h \leq p$. P-M & Wolf (2021)

Intuition:

1) $\hat{\theta}_h^{\text{VAR}} = \text{MSE-optimal forecast given model-implied autocov's out to horizon } h.$

2 VAR(p) estimator consistently estimates true autocov's Cov($Y_t, Y_{t-\ell}$) for $\ell \leq p$.

- **3** LP estimates the MSE-optimal forecast by definition.
- Note: Result does not assume any particular parametric model (e.g., linear or VAR(p)).

Implications of large-sample equivalence

- **1** LP \approx VAR at short horizons *h*, but not necessarily for h > p.
- **2** LP IRF \approx VAR IRF with sufficiently large *p*.
 - In fact, estimators are asymptotically equivalent as $p, T
 ightarrow \infty$. Xu (2022)
- **3** LP and VAR are not conceptually distinct paradigms. Just two techniques for estimating projections with shared large-sample estimand.
- **4** Identification \perp estimation: Any SVAR identification procedure can be implemented using LP and *vice versa* (also when x_t is not observed).
 - Short-run/long-run zero restrictions, sign restrictions, narrative restrictions, higher moments.
- **5** LP is not more "robust to misspecification" than VAR (unless this means sensitivity to p).

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Finite-sample bias-variance trade-off

- At horizons h > p, we face non-trivial bias-variance trade-off:
 - VAR extrapolates longer-run responses from first *p* autocov's. Low variance, potentially high bias if DGP ≠ VAR(*p*).
 - LP does not extrapolate. High variance, low bias.
- In ongoing simulation project, my coauthors and I explore this trade-off across 1,000s of empirically calibrated DGPs (caveat: stationarity). Li, P-M & Wolf (2022)



Medians across 6,000 DGPs. Source: Li, P-M & Wolf (2022).

Improving the trade-off through shrinkage/penalization

- Shrinkage/penalization: nudge least-squares LP/VAR estimates towards *a priori* more reasonable values.
 - Bayesian VAR: shrink towards impulse responses for independent random walks or white noise. Doan, Litterman & Sims (1984)
 - Smoothness: shrink LP towards impulse responses that are smoother functions of *h*. Shiller (1973); P-M (2016); Barnichon & Brownlees (2019)
 - Functional form: shrink LP towards VAR-implied impulse responses or towards exponential shapes. Barnichon & Matthes (2018); Miranda-Aggripino & Ricco (2021)
- Introduces some bias (unless prior is exactly right) in order to lower variance.
- Degree of shrinkage chosen based on Bayesian considerations or frequentist objective fct.



Fraction of DGPs where estimators are preferred according to loss function = $\omega \times bias^2 + (1 - \omega) \times variance$. Source: Li, P-M & Wolf (2022).

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Confidence intervals for impulse responses

- VAR(*p*)-based CIs are non-robust at longer horizons (lack uniform validity).
 - Consider AR(1): $\hat{\theta}_h = \hat{\rho}^h$.
 - Sampling distr'n highly sensitive to estimation error ρ̂ − ρ when h → ∞. Methods that are tailored to the case h → ∞ don't work at shorter horizons.
 Wright (2000); Gospodinov (2004); Pesavento and Rossi (2007), Mikusheva (2012)
 - Unit roots can cause non-normal limiting distr'n (not always). Want to avoid pre-testing for unit roots. Inoue & Kilian (2002, 2020)
- In contrast, linearity of LP means normal approx'n works regardless of *h* and unit roots. Dufour, Pelletier & Renault (2006); Breitung & Brüggemann (2019); Montiel Olea & P-M (2021)
 - Important to control for lagged data so variance of OLS residual doesn't blow up.
 - Bonus: No need for Newey-West standard errors if we're regressing on a shock x_t .

Can we shorten the CIs?

- LP (or VAR with large p) often yields wide CIs in practice. Can shrinkage help here?
 - <u>No</u>, not if we insist on usual coverage notion: 1α coverage prob. separately at each horizon, regardless of true impulse responses. Pratt (1961); Amstrong & Kolesár (2018)
 - Intuition: Must guard against worst-case bias over parameter space.
- If we want narrower CIs, we have to relax the coverage requirement. Ideas:
 - Control coverage prob. only on average across horizons. Armstrong, Kolesár & P-M (2022)
 - Restrict parameter space (e.g., *a priori* upper bound on jaggedness of IRF).
 - Coverage for simpler "surrogate IRF". Genovese & Wasserman (2008)
 - Don't report confidence band. Think harder about what specific hypothesis we want to test.

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Impulse responses in nonlinear models

- What do linear LP/VAR estimate in nonlinear DGPs?
- Assume general nonlinear causal representation, given shock vector $\varepsilon_t = (x_t, \varepsilon_{2:n,t})$:

$$y_{t+h} = g_h(\varepsilon_{t+h},\ldots,\varepsilon_{t+1},\varepsilon_t,Y_{t-1}).$$

- E.g., binary outcomes, ZLB, discrete/smooth regime-switching, higher-order lag dynamics.
- Many possible definitions of nonlinear impulse responses. Linear LP/VAR do not consistently estimate all of these. Gonçalves, Herrera, Kilian & Pesavento (2021, 2022)
- Consider manipulating x_t from δ_1 to δ_2 , while averaging over other shocks and the history:

$$\theta_h(\delta_1,\delta_2) \equiv E\Big[g_h(\varepsilon_{t+h},\ldots,\varepsilon_{t+1},(\delta_2,\varepsilon_{2:n,t}),Y_{t-1}) - g_h(\varepsilon_{t+h},\ldots,\varepsilon_{t+1},(\delta_1,\varepsilon_{2:n,t}),Y_{t-1})\Big].$$

Interpretation of linear LP/VAR estimand in nonlinear models

• If shocks are mutually and serially independent, and $h \leq p$,

$$\lim_{T \to \infty} \hat{\theta}_h^{\mathsf{LP}} = \lim_{T \to \infty} \hat{\theta}_h^{\mathsf{VAR}} = \lim_{\delta \to 0} \int_{-\infty}^{\infty} w(x) \frac{\theta_h(x, x + \delta)}{\delta} \, dx,$$

where $w(\cdot) \ge 0$ and $\int_{-\infty}^{\infty} w(x) dx = 1$. Yitzhaki (1996); Angrist & Pischke (2009); Rambachan & Shephard (2021); Kolesár & P-M (2022)

- Linear LP/VAR estimate a meaningful summary: weighted average of scaled causal effects $\theta_h(x, x + \delta)/\delta$ for infinitesimal shock size $\delta \approx 0$.
- Weight function $w(\cdot)$ does not depend on *h*. Estimable. Angrist & Krueger (1999)
- Ongoing work: implications for state/sign-dependence, compare with nonlinear estim'rs.

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Conclusions and areas for future research

- LP and VAR belong to a menu of procedures with shared large-sample estimand.
- Shrinkage/penalization usually improves bias-variance trade-off.
 - Can new procedures further optimize trade-off? Does trade-off differ when we have panel data?
- However, shrinkage does not help confidence interval construction in the usual sense.
 - Should we relax the coverage requirement, and how? How to deal with very long horizons? Müller & Watson (2017, 2018)
- Linear LP/VAR useful also in nonlinear DGPs, but much work to be done.
 - Do existing nonlinear estimators have interpretable estimands under misspecification? Can we hope to accurately estimate specific nonlinearities in macro data, and how?